

LOGARITHMIC SOBOLEV INEQUALITIES IN DISCRETE PRODUCT SPACES: (Marton, 2015)

Let \mathcal{X} be a finite set, and \mathcal{X}^n be the set of n -sequences from \mathcal{X} .

Let $\mathcal{P}(\mathcal{X}^n)$ be the probability simplex on \mathcal{X}^n .

Def: (W_2 -distance) Suppose $X^n \sim P_{X^n} \in \mathcal{P}(\mathcal{X}^n)$ and $Y^n \sim P_{Y^n} \in \mathcal{P}(\mathcal{X}^n)$. Then, we define:

$$W_2(P_{X^n}, P_{Y^n}) \triangleq \min_{P_{X^n, Y^n}} \sqrt{\sum_{i=1}^n P_{X^n, Y^n}(X_i \neq Y_i)^2}$$

where the minimum is over all couplings of P_{X^n} and P_{Y^n} .

Remark: • This is not a 2-Wasserstein distance wrt Hamming metric.

- Minimum is achieved.
- W_2 is a metric on $\mathcal{P}(\mathcal{X}^n)$.

Theorem 1: (W_2^2 tensorization \Rightarrow KL tensorization)

Fix a distribution $Q_{X^n} \in \mathcal{P}(\mathcal{X}^n)$, and let $\alpha \triangleq \min_{\substack{1 \leq i \leq n \\ x^i \in \mathcal{X}^n: Q_{X^n}(x^i) > 0}} Q_{X_i | X_{-i}}(x_i | x_{-i})$, where for any string

x^n and any set $S \subseteq [n] \triangleq \{1, \dots, n\}$, $x_S \triangleq \{x_i : i \in S\}$, and $x_i = x_{\{i\}}$ and $x_{-i} = x_{[n] \setminus \{i\}}$.

Fix a distribution $P_{Y^n} \in \mathcal{P}(\mathcal{X}^n)$ such that $P_{Y^n} \ll Q_{X^n}$. absolutely continuous

Suppose for all $I \subseteq [n]$ and all $y_{I^c} \in \mathcal{X}^{[n] \setminus I}$, we have:

$$W_2^2(P_{Y_I | Y_{I^c}}(\cdot | y_{I^c}), Q_{X_I | X_{I^c}}(\cdot | y_{I^c})) \leq C \sum_{i \in I} \underbrace{\mathbb{E} \left[\left\| P_{Y_i | Y_{-i}}(\cdot | y_{-i}) - Q_{X_i | X_{-i}}(\cdot | y_{-i}) \right\|_{TV}^2 \right]}_{\substack{\text{wrt } P_{Y_i | Y_{-i}} = q_i \\ \uparrow \text{conditional } W_2^2\text{-distance}}} \Big| Y_{I^c} = y_{I^c}$$

for a universal constant C . Then, we have:

$$D(P_{Y^n} \| Q_{X^n}) \leq \underbrace{\frac{4C}{\alpha}}_{\substack{\text{reverse} \\ \text{Pinsker}}} \sum_{i=1}^n \underbrace{\mathbb{E} \left[\left\| P_{Y_i | Y_{-i}}(\cdot | y_{-i}) - Q_{X_i | X_{-i}}(\cdot | y_{-i}) \right\|_{TV}^2 \right]}_{\substack{\text{wrt } P_{Y_i} \\ \uparrow \text{Pinsker}}} \leq \frac{2C}{\alpha} \sum_{i=1}^n \underbrace{D(P_{Y_i | Y_{-i}} \| Q_{X_i | X_{-i}} | P_{Y_{-i}})}_{\substack{\uparrow \text{conditional KL divergence}}}$$

Proof: We prove this by induction. Assume the conditions of the theorem hold, and that we have proven it for $n-1$:

$$D(P_{Y_i | Y_{-i}}(\cdot | y_{-i}) \| Q_{X_i | X_{-i}}(\cdot | y_{-i})) \leq \frac{4C}{\alpha} \sum_{j \neq i} \mathbb{E} \left[\left\| P_{Y_j | Y_{-j}}(\cdot | y_{-j}) - Q_{X_j | X_{-j}}(\cdot | y_{-j}) \right\|_{TV}^2 \Big| Y_i = y_i \right] \leftarrow \text{inductive hypothesis}$$

for all $i \in [n]$ and all $y_i \in \mathcal{X}$.

First, by averaging chain rules, we get:

$$D(P_{Y^n} \| Q_{X^n}) = \underbrace{\frac{1}{n} \sum_{i=1}^n D(P_{Y_i} \| Q_{X_i})}_{\textcircled{1}} + \underbrace{\frac{1}{n} \sum_{i=1}^n D(P_{Y_i | Y_{-i}} \| Q_{X_i | X_{-i}} | P_{Y_{-i}})}_{\textcircled{2}} \quad [*]$$

[continued]

Proof continued:

To bound ②, we use the inductive hypothesis to get:

$$\begin{aligned} \frac{1}{n} \sum_{i=1}^n D(P_{Y_i|\cdot|Y_i} \| Q_{X_i|\cdot|X_i} | P_{Y_i}) &\leq \frac{1}{n} \frac{4C}{\alpha} \sum_{i=1}^n \sum_{j \neq i} \mathbb{E} \left[\left\| P_{Y_j|Y_{j^c}(\cdot|Y_{j^c})} - Q_{X_j|X_{j^c}(\cdot|Y_{j^c})} \right\|_{TV}^2 \right] \\ &\quad \uparrow \text{does not depend on } i \Rightarrow (n-1) \text{ copies for each } j \\ &= \left(1 - \frac{1}{n}\right) \frac{4C}{\alpha} \sum_{j=1}^n \mathbb{E} \left[\left\| P_{Y_j|Y_{j^c}(\cdot|Y_{j^c})} - Q_{X_j|X_{j^c}(\cdot|Y_{j^c})} \right\|_{TV}^2 \right]. \quad [*] \end{aligned}$$

To bound ①, notice from reverse Pinsker's inequality ($D(P\|Q) \leq \frac{4}{\min(Q)} \|P-Q\|_{TV}^2$ for $Q > 0$):

$$D(P_{Y_i} \| Q_{X_i}) \leq \frac{4}{\alpha} \|P_{Y_i} - Q_{X_i}\|_{TV}^2$$

for all $i \in [n]$, where we use the fact that: $\forall x \in \mathcal{X}, \forall i \in [n], Q_{X_i}(x) \geq \alpha$.

$$\left[Q_{X_i}(x) = \sum_{x_i^c} \frac{Q_{X_i|X_{i^c}(x|x_i^c)} Q_{X_{i^c}}(x_i^c)}{\geq \alpha \text{ if } Q_{X_i|X_{i^c}(x,x_i^c)} > 0} \right]$$

Let π be the optimal W_2 -coupling of P_{Y^n}, Q_{X^n} . Then, we have:

$$W_2^2(P_{Y^n}, Q_{X^n}) = \sum_{i=1}^n \mathbb{P}_{\pi}(X_i \neq Y_i)^2 \geq \sum_{i=1}^n \|P_{Y_i} - Q_{X_i}\|_{TV}^2 \quad [\text{maximal coupling of TV}]$$

which implies using the condition in the theorem with $I = [n]$ that:

$$\begin{aligned} \frac{1}{n} \sum_{i=1}^n D(P_{Y_i} \| Q_{X_i}) &\leq \frac{1}{n} \frac{4}{\alpha} \sum_{i=1}^n \|P_{Y_i} - Q_{X_i}\|_{TV}^2 \leq \frac{1}{n} \frac{4}{\alpha} W_2^2(P_{Y^n}, Q_{X^n}) \\ &\leq \frac{1}{n} \frac{4C}{\alpha} \sum_{i=1}^n \mathbb{E} \left[\left\| P_{Y_i|Y_{i^c}(\cdot|Y_{i^c})} - Q_{X_i|X_{i^c}(\cdot|Y_{i^c})} \right\|_{TV}^2 \right]. \quad [*] \end{aligned}$$

Hence, we get:

$$\begin{aligned} D(P_{Y^n} \| Q_{X^n}) &\leq \frac{4C}{\alpha} \sum_{i=1}^n \mathbb{E} \left[\left\| P_{Y_i|Y_{i^c}(\cdot|Y_{i^c})} - Q_{X_i|X_{i^c}(\cdot|Y_{i^c})} \right\|_{TV}^2 \right] \quad (\text{using } [*]) \\ &\leq \frac{2C}{\alpha} \sum_{i=1}^n D(P_{Y_i|Y_{i^c}} \| Q_{X_i|X_{i^c}} | P_{Y_i}) \end{aligned}$$

where the final inequality follows from Pinsker's inequality:

$$D(P\|Q) \geq 2 \|P-Q\|_{TV}^2.$$



Remark: Doing the induction from $k-1$ to k would require us to use more conditional expectations, and would illustrate the full power of the condition in the theorem.

Idea: W_2 -distance related to TV distance, KL distance bounded by TV distance (Pinsker + reverse Pinsker), use this to prove W_2^2 -tensorization \Rightarrow KL tensorization.
($KL \leq TV^2 \leq W_2^2 \leq TV^2 \leq KL$)

[continued.]

Def: (Gibbs Sampler) For $i \in [n]$, define the Markov chain:

$$T_i(x^n | y^n) \triangleq \mathbb{1}\{x_i = y_i\} Q_{x_i | x_{-i}}(x_i | y_i), \quad \forall x^n, y^n \in \mathcal{X}^n$$

for a distribution $Q_{X^n} \in \mathcal{P}(\mathcal{X}^n)$. The Gibbs sampler for Q_{X^n} is the Markov chain:

$$\Gamma \triangleq \frac{1}{n} \sum_{i=1}^n T_i$$

↑ row stochastic matrices

Observe that for any $x^n \in \mathcal{X}^n$:

$$\begin{aligned} Q_{X^n} T_i(x^n) &= \sum_{y^n} Q_{X^n}(y^n) T_i(x^n | y^n) = \sum_{y^n} Q_{X^n}(y^n) \mathbb{1}\{x_i = y_i\} Q_{x_i | x_{-i}}(x_i | y_i) \\ &= Q_{x_i | x_{-i}}(x_i | x_{-i}) \sum_{y_i} Q_{x_i, x_{-i}}(y_i, x_{-i}) = Q_{x_i | x_{-i}}(x_i | x_{-i}) Q_{x_i}(x_i) \\ &= Q_{X^n}(x^n) \end{aligned}$$

which means Q_{X^n} is the invariant measure for T_i and Γ .

The Dirichlet form associated with Γ is:

$$\mathcal{E}_\Gamma(f, f) \triangleq \left\langle \underbrace{(I - \Gamma)}_{\text{identity map}} f, f \right\rangle_{Q_{X^n}} = \sum_{x^n} (f(x^n) - \Gamma f(x^n)) f(x^n) Q_{X^n}(x^n) = \frac{1}{2} \sum_{x^n, y^n} T(y^n | x^n) Q_{X^n}(x^n) (f(x^n) - f(y^n))^2$$

for all $f \in \mathcal{L}^2(\mathcal{X}^n, Q_{X^n})$. The logarithmic Sobolev inequality for Γ is:

$$D(P_{Y^n} \| Q_{X^n}) \leq c \mathcal{E}_\Gamma \left(\sqrt{\frac{f}{P_{Y^n}}}, \sqrt{\frac{f}{Q_{X^n}}} \right) \quad \text{for all } P_{Y^n} \in \mathcal{P}(\mathcal{X}^n)$$

where the best constant c is called the log-Sobolev constant.

Remark: In some papers, $\frac{1}{c}$ is the log-Sobolev constant.

Corollary: (Gibbs Sampler contraction)

If $Q_{X^n}, P_{Y^n} \in \mathcal{P}(\mathcal{X}^n)$ satisfy the conditions of Theorem 1, then:

[KL Contraction] ① $D(P_{Y^n} \Gamma \| Q_{X^n}) \leq \left(1 - \frac{\alpha}{2nC}\right) D(P_{Y^n} \| Q_{X^n})$, { Hold for all P_{Y^n} when Q_{X^n} satisfies Dobrushin uniqueness condition.

[Log-Sobolev] ② $D(P_{Y^n} \| Q_{X^n}) \leq \frac{4Cn}{\alpha} \mathcal{E}_\Gamma \left(\sqrt{\frac{P_{Y^n}}{Q_{X^n}}}, \sqrt{\frac{P_{Y^n}}{Q_{X^n}}} \right)$.

Proof: ① First observe that:

$$\begin{aligned} D(P_{Y^n} \| Q_{X^n}) &= D(P_{Y_i} \| Q_{X_i}) + D(P_{Y_i | Y_{-i}} \| Q_{X_i | X_{-i}} | P_{Y_{-i}}) \quad [\text{Chain rule}] \\ &= D(P_{Y_i} \Gamma_i \| Q_{X_i}) + D(P_{Y_i | Y_{-i}} \| Q_{X_i | X_{-i}} | P_{Y_{-i}}) \\ &= \frac{1}{n} \sum_{i=1}^n D(P_{Y_i} \Gamma_i \| Q_{X_i}) + \frac{1}{n} \sum_{i=1}^n D(P_{Y_i | Y_{-i}} \| Q_{X_i | X_{-i}} | P_{Y_{-i}}) \end{aligned}$$

By convexity of KL divergence:

$$D(P_{Y^n} \Gamma \| Q_{X^n}) \leq \frac{1}{n} \sum_{i=1}^n D(P_{Y_i} \Gamma_i \| Q_{X_i}) \leq D(P_{Y^n} \| Q_{X^n}) - \frac{\alpha}{2nC} D(P_{Y^n} \| Q_{X^n})$$

using the KL tensorization result of Theorem 1.

Theorem 2: (Dobrushin's condition $\Rightarrow W_2^2$ tensorization)

If $Q_{X^n} \in \mathcal{P}(X^n)$ satisfies Dobrushin's uniqueness condition with coupling matrix A with $\|A\|_2 < 1$, then for any $P_{Y^n} \in \mathcal{P}(X^n)$:

conds of Thm 1 $\left\{ \begin{aligned} W_2^2(P_{Y_I|Y_{I^c}}(\cdot|y_{I^c}), Q_{X_I|X_{I^c}}(\cdot|y_{I^c})) &\leq C \sum_{i \in I} \mathbb{E} \left[\left\| P_{X_i|X_{I^c}}(\cdot|y_{I^c}) - Q_{X_i|X_{I^c}}(\cdot|y_{I^c}) \right\|_{TV}^2 \mid Y_{I^c} = y_{I^c} \right] \\ \text{for all } I \subseteq [n] \text{ and all } y_{I^c} \in X^{[n] \setminus I}, \text{ where } C &= \frac{1}{(1 - \|A\|_2)^2}. \end{aligned} \right.$

Proof: It suffices to prove this for $I = [n]$. (For any $I \subseteq [n]$ and any $y_{I^c} \in X^{[n] \setminus I}$, $Q_{X_I|X_{I^c}}(\cdot|y_{I^c})$ satisfies Dobrushin's uniqueness condition with a minor of A , $\{a_{k,i}\}_{k,i \in I} = A_I$, as coupling matrix. Moreover, $A_I^e = D_1 A D_2 \Rightarrow \|A_I\|_2 = \|A_I^e\|_2 = \|D_1 A D_2\|_2 \leq \underbrace{\|D_1\|_2}_{\leq 1} \|A\|_2 \underbrace{\|D_2\|_2}_{\leq 1} \leq \|A\|_2$ extend A_I with zero rows and columns so it is $n \times n$ diagonal matrices with 1's at I and 0's elsewhere. So, $C = 1/(1 - \|A\|_2)^2$ is a valid constant for $I \subseteq [n]$.)

which implies that $\frac{1}{(1 - \|A\|_2)^2} \geq \frac{1}{(1 - \|A_I\|_2)^2}$.

Idea: Prove that Gibbs sampler T is a contraction wrt W_2 -distance, and then use relationship between W_2 and TV distances.

Fix two distributions $R_{U^n}, S_{Z^n} \in \mathcal{P}(X^n)$. Let π be a coupling between them.

Construct random variables $(U'_1, \dots, U'_n), (Z'_1, \dots, Z'_n)$ as follows:

- Sample a random variable $t \sim \text{Unif}[n]$ to get $t = i$, (t indep. of everything else)
- Let $U'_k = U_k$ and $Z'_k = Z_k$ for $k \neq i$ so that the distribution of U'_i, Z'_i given $t = i$ is $\pi_{U'_i, Z'_i}$.
- Let the distribution of U'_i, Z'_i given $t = i$ and $U'_i = u_{i,c}, Z'_i = z_{i,c}$ be the maximal coupling of $Q_{X_i|X_{i^c}}(\cdot|u_{i,c})$ and $Q_{X_i|X_{i^c}}(\cdot|z_{i,c})$ that achieves $\|Q_{X_i|X_{i^c}}(\cdot|u_{i,c}) - Q_{X_i|X_{i^c}}(\cdot|z_{i,c})\|_{TV}$.
 \uparrow dist. of U'_i given $U'_i = u_{i,c}, t = i$ \uparrow dist. of Z'_i given $Z'_i = z_{i,c}, t = i$

Clearly, the distribution of U^n is $R_{U^n} T$ and the distribution of Z^n is $S_{Z^n} T$.

For any $j \in [n]$,

$$\begin{aligned} P(U'_j \neq Z'_j) &= P(t \neq j) P(U'_j \neq Z'_j | t \neq j) + P(t = j) P(U'_j \neq Z'_j | t = j) \\ &= (1 - \frac{1}{n}) P(U_j \neq Z_j) + \frac{1}{n} \sum_{u_{j,c}, z_{j,c}} P(U'_j \neq Z'_j | t = j, U'_j = u_{j,c}, Z'_j = z_{j,c}) P(U'_j = u_{j,c}, Z'_j = z_{j,c} | t = j) \\ &= (1 - \frac{1}{n}) P(U_j \neq Z_j) + \frac{1}{n} \sum_{u_{j,c}, z_{j,c}} \underbrace{\|Q_{X_j|X_{j^c}}(\cdot|u_{j,c}) - Q_{X_j|X_{j^c}}(\cdot|z_{j,c})\|_{TV}}_{= \pi_{U'_j, Z'_j}(u_{j,c}, z_{j,c})} P(U'_j = u_{j,c}, Z'_j = z_{j,c} | t = j) \\ &= (1 - \frac{1}{n}) P(U_j \neq Z_j) + \frac{1}{n} \mathbb{E} \left[\sum_{k \neq j} \mathbb{1}\{U_k \neq Z_k\} a_{k,j} \right] \leftarrow \text{use triangle inequality and Dobrushin's condition} \\ &= (1 - \frac{1}{n}) P(U_j \neq Z_j) + \frac{1}{n} \sum_{k \neq j} a_{k,j} P(U_k \neq Z_k). \end{aligned}$$

Proof continued:

Let $p = [P(U_1 \neq Z_1) \dots P(U_n \neq Z_n)]$ and $p' = [P(U'_1 \neq Z'_1) \dots P(U'_n \neq Z'_n)]$.

Then, we have:

$$P' \leq P \left[\left(1 - \frac{1}{n}\right) I + \frac{1}{n} A \right] \Rightarrow \|P'\|_2 \leq \|P\|_2 \left\| \left(1 - \frac{1}{n}\right) I + \frac{1}{n} A \right\|_2 \leq \left(1 - \frac{1}{n}(1 - \|A\|_2)\right) \|P\|_2$$

entrywise identity matrix diagonal entries = 0 definition of $\|\cdot\|_2$ for matrix Δ -inequality

$$\Rightarrow \sqrt{\sum_{j=1}^n P(U'_j \neq Z'_j)^2} \leq \left(1 - \frac{1}{n}(1 - \|A\|_2)\right) \sqrt{\sum_{j=1}^n P(U_j \neq Z_j)^2}$$

wrt π

Choosing the minimum W_2 -couplings, we get:

$$\underline{W_2(R_{U^n T}, S_{Z^n T})} \leq \left(1 - \frac{1}{n}(1 - \|A\|_2)\right) W_2(R_{U^n}, S_{Z^n}). \quad [\text{contraction of } T \text{ in } W_2\text{-distance}]$$

By triangle inequality, for any $R_{Y^n} \in \mathcal{P}(X^n)$: $= Q_{X^n T}$ [invariant measure]

$$\begin{aligned} W_2(R_{Y^n}, Q_{X^n}) &\leq W_2(R_{Y^n}, R_{Y^n T}) + W_2(R_{Y^n T}, \overline{Q_{X^n}}) \\ &\leq W_2(R_{Y^n}, R_{Y^n T}) + \left(1 - \frac{1}{n}(1 - \|A\|_2)\right) W_2(R_{Y^n}, Q_{X^n}) \quad [\text{contraction}] \end{aligned}$$

$$\Rightarrow W_2(R_{Y^n}, Q_{X^n}) \leq \frac{n}{1 - \|A\|_2} W_2(R_{Y^n}, R_{Y^n T}). \quad [*]$$

Now we upper bound $W_2(R_{Y^n}, R_{Y^n T})$. We construct a coupling of $R_{Y^n}, R_{Y^n T} \in \mathcal{P}(X^n)$ using random variables Y^n, Y'^n :

- ① Sample a random variable $t \sim \text{Unif}[n]$ to get $t=i$, (t indep. of everything else)
 - ② Let $Y_k = Y'_k$ for $k \neq i$ so that the distribution of $Y_{i \cdot}$ given $t=i$ is $P_{Y_{i \cdot}}$,
 - ③ Let the distribution of Y_i, Y'_i given $t=i$ and $Y_{i \cdot} = y_{i \cdot}$ be the maximal coupling of $P_{Y_i | Y_{i \cdot}}(\cdot | y_{i \cdot})$ and $Q_{Y_i | Y_{i \cdot}}(\cdot | y_{i \cdot})$ that achieves $\|P_{Y_i | Y_{i \cdot}}(\cdot | y_{i \cdot}) - Q_{Y_i | Y_{i \cdot}}(\cdot | y_{i \cdot})\|_{TV}$.
- dist. of Y_i given $t=i$ and $Y_{i \cdot} = y_{i \cdot}$ dist. of Y'_i given $t=i$ and $Y_{i \cdot} = y_{i \cdot}$

Clearly, the distribution of Y^n is R_{Y^n} and the distribution of Y'^n is $R_{Y^n T}$.

For any $j \in [n]$,

$$\begin{aligned} P(Y_j \neq Y'_j) &= \frac{1}{n} P(Y_j \neq Y'_j | t=j) + \underbrace{\left(1 - \frac{1}{n}\right) P(Y_j \neq Y'_j | t \neq j)}_{=0} \\ &= \frac{1}{n} \sum_{y_{j \cdot}} P(Y_j \neq Y'_j | t=j, Y_{j \cdot} = y_{j \cdot}) P(Y_{j \cdot} = y_{j \cdot} | t=j) \\ &= \frac{1}{n} \mathbb{E}_{P_{Y_{j \cdot}}} \left[\|P_{Y_j | Y_{j \cdot}}(\cdot | Y_{j \cdot}) - Q_{Y_j | Y_{j \cdot}}(\cdot | Y_{j \cdot})\|_{TV} \right] \end{aligned}$$

$$\Rightarrow P(Y_j \neq Y'_j)^2 = \frac{1}{n^2} \mathbb{E} \left[\|P_{Y_j | Y_{j \cdot}}(\cdot | Y_{j \cdot}) - Q_{Y_j | Y_{j \cdot}}(\cdot | Y_{j \cdot})\|_{TV}^2 \right] \leq \frac{1}{n^2} \mathbb{E} \left[\|P_{Y_j | Y_{j \cdot}}(\cdot | Y_{j \cdot}) - Q_{Y_j | Y_{j \cdot}}(\cdot | Y_{j \cdot})\|_{TV}^2 \right].$$

Jensen's inequality

$$\text{Hence, we have: } W_2(R_{Y^n}, R_{Y^n T}) \leq \sqrt{\sum_{j=1}^n P(Y_j \neq Y'_j)^2} \leq \frac{1}{n} \sqrt{\sum_{j=1}^n \mathbb{E} \left[\|P_{Y_j | Y_{j \cdot}}(\cdot | Y_{j \cdot}) - Q_{Y_j | Y_{j \cdot}}(\cdot | Y_{j \cdot})\|_{TV}^2 \right]},$$

which with [*] completes the proof. ▣